

2.3: The correlation coefficient .

Recall :

$$\rho_{xy} = \frac{\sigma_{xy}}{\sigma_x \sigma_y}$$

$$\sigma_y = \sqrt{\frac{1}{N} \sum (y_i - \mu_y)^2} \quad , \quad \sigma_x = \sqrt{\frac{1}{N} \sum (x_i - \mu_x)^2}$$

$$\sigma_{xy} = \frac{1}{N} \sum (x_i - \mu_x)(y_i - \mu_y)$$

$$-1 \leq \rho \leq 1$$

Data set : $(x_1, y_1), \dots, (x_N, y_N)$.

Def: let x_1, x_2 be Random variable with joint p.d.f $f(x_1, x_2)$. We define the covariance of x_1 and x_2 as follows :

$$\text{cov}(x_1, x_2) = E \left[(x_1 - \mu_1)(x_2 - \mu_2) \right] \text{ where } \mu_1 = E(x_1) \quad , \quad \mu_2 = E(x_2) .$$

proposition : $\text{cov}(x_1, x_2) = E(x_1 x_2) - \mu_1 \mu_2$.

Proof :

$$\begin{aligned} \text{cov}(x_1, x_2) &= E \left((x_1 - \mu_1)(x_2 - \mu_2) \right) \\ &= E \left(x_1 x_2 - x_1 \mu_2 - x_2 \mu_1 + \mu_1 \mu_2 \right) \\ &= E(x_1 x_2) - \mu_1 \mu_2 - \mu_2 \mu_1 + \mu_1 \mu_2 \\ &= E(x_1 x_2) - \mu_1 \mu_2 \end{aligned}$$

□

Def: let X_1, X_2 be Random variable with joint p.d.f $f(x_1, x_2)$. we define the correlation coefficient of X_1 and X_2 as follows:

$$\rho_{xy} = \frac{\text{cov}(X_1, X_2)}{\sigma_1 \sigma_2} \quad \text{where } \sigma_1 = \sqrt{\text{var}(X_1)}, \quad \sigma_2 = \sqrt{\text{var}(X_2)}.$$

RMK: $\text{cov}(X_1, X_2)$ assume $E(X_1), E(X_2)$ exist.

ρ assume $\sqrt{\text{var}(X_1)} > 0, \sqrt{\text{var}(X_2)} > 0$.

Exp 1: X, y are Random variable with joint p.d.f,
$$f(x, y) = \begin{cases} x+y, & 0 \leq x < 1, 0 \leq y < 1 \\ 0, & \text{elsewhere} \end{cases}$$
 Find ρ .

$$\rho = \frac{\text{cov}(X, y)}{\sigma_x \sigma_y} = \frac{E(X - \mu_x)(y - \mu_y)}{\sigma_x \sigma_y} = \frac{E(Xy) - \mu_x \mu_y}{\sigma_x \sigma_y}$$

$$\rightarrow \mu_x = E(X) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x f(x, y) dx dy = \int_0^1 \int_0^1 x(x+y) dx dy = \int_0^1 \int_0^1 x^2 + xy dx dy$$

$$= \int_0^1 \left[\frac{x^3}{3} + \frac{x^2}{2} y \right]_0^1 dy$$

$$= \int_0^1 \left[\frac{1}{3} + \frac{1}{2} y \right] dy$$

$$= \left[\frac{1}{3} y + \frac{1}{2} \frac{y^2}{2} \right]_0^1$$

$$= \frac{1}{3} + \frac{1}{4}$$

$$= \frac{7}{12}$$

$$\begin{aligned}
 \Rightarrow \mu_y = E(y) &= \iint_{-\infty}^{\infty} y f(x,y) dx dy = \int_0^1 \int_0^1 y(x+y) dx dy = \int_0^1 \int_0^1 xy + y^2 dx dy \\
 &= \int_0^1 \left[\frac{x^2}{2} y + y^2 x \right]_0^1 dy \\
 &= \int_0^1 \left[\frac{1}{2} y + y^2 \right] dy \\
 &= \left[\frac{1}{2} \frac{y^2}{2} + \frac{y^3}{3} \right]_0^1 = \frac{7}{12}
 \end{aligned}$$

$$\begin{aligned}
 \Rightarrow \sigma_x^2 = \text{var}(x) &= E(x - \mu_x)^2 = E(x^2) - \mu_x^2 \\
 &= \iint_{-\infty}^{\infty} x^2 f(x,y) dx dy - \mu_x^2 \\
 &= \int_0^1 \int_0^1 x^3 + x^2 y dx dy - \left(\frac{7}{12} \right)^2 \\
 &= \int_0^1 \left[\frac{x^4}{4} + \frac{x^3}{3} y \right]_0^1 dy - \frac{49}{144} \\
 &= \int_0^1 \left[\frac{1}{4} y + \frac{1}{3} \frac{y^2}{2} \right] dy - \frac{49}{144} \\
 &= \left[\frac{36 \times 1}{36 \times 4} + \frac{24 \times 1}{24 \times 6} \right] - \frac{49}{144} \\
 &= \frac{36 + 24 - 49}{144} \\
 &= \frac{11}{144}
 \end{aligned}$$

$$\begin{aligned}
 \Rightarrow \sigma_y^2 &= \text{Var}(y) = E(y - \mu_y)^2 = E(y^2) - \mu_y^2 \\
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} y^2 f(x, y) dx dy - \mu_y^2 \\
 &= \int_0^1 \int_0^1 y^2 x + y^3 dx dy - \left(\frac{7}{12}\right)^2 \\
 &= \frac{11}{144}
 \end{aligned}$$

$$\begin{aligned}
 \Rightarrow E(xy) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy f(x, y) dx dy = \int_0^1 \int_0^1 xy (x+y) dx dy \\
 &= \int_0^1 \left. \frac{x^3}{3} y + \frac{x^2}{2} y^2 \right|_0^1 dy \\
 &= \int_0^1 \left(\frac{1}{3} y + \frac{1}{2} y^2 \right) dy \\
 &= \left. \frac{1}{3} \frac{y^2}{2} + \frac{1}{2} \frac{y^3}{3} \right|_0^1 \\
 &= \frac{1}{6} + \frac{1}{6} = \frac{1}{3}
 \end{aligned}$$

$$\begin{aligned}
 \Rightarrow \rho &= \frac{E(xy) - \mu_x \mu_y}{\sigma_x \sigma_y} = \frac{\frac{1}{3} - \left(\frac{7}{12}\right)\left(\frac{7}{12}\right)}{\sqrt{\frac{11}{144}} \sqrt{\frac{11}{144}}} \\
 &= \frac{-1}{11} \\
 &= -0.09
 \end{aligned}$$

Remark: Assume x and y satisfy $E(y|x) = a + bx$, we can show:

$$a = \mu_2 - \rho \frac{\sigma_2}{\sigma_1} \mu_1 \quad \text{and} \quad b = \rho \frac{\sigma_2}{\sigma_1}$$

$$\text{Hence, } E(y|x) = \mu_2 - \rho \frac{\sigma_2}{\sigma_1} \mu_1 + \rho \frac{\sigma_2}{\sigma_1} x$$

$$\Rightarrow E(y|x) = \mu_2 + \rho \frac{\sigma_2}{\sigma_1} (x - \mu_1)$$

where:

Similarly if $E(x|y) = c + dy$

$$\mu_1 = \mu_x = E(x)$$

$$\Rightarrow E(x|y) = \mu_1 + \rho \frac{\sigma_1}{\sigma_2} (y - \mu_2)$$

$$\mu_2 = \mu_y = E(y)$$

$$\sigma_1 = \sqrt{\sigma_1^2} = \sqrt{\sigma_x^2} = \sqrt{\text{Var}(x)}$$

$$\sigma_2 = \sqrt{\sigma_2^2} = \sqrt{\sigma_y^2} = \sqrt{\text{Var}(y)}$$

and we can show:

$$E(\text{var}(y|x)) = \sigma_2^2 (1 - \rho^2)$$

$$\rho = \frac{\text{Cov}(x,y)}{\sigma_1 \sigma_2}$$

Remark: $E(\text{var}(y|x)) \geq 0$

$$\Rightarrow \sigma_2^2 (1 - \rho^2) \geq 0$$

$$\Rightarrow 1 - \rho^2 \geq 0$$

$$\Rightarrow -1 \leq \rho \leq 1$$

example 2: $E(y|x) = 4x + 3 = a + bx$

$$E(x|y) = \frac{1}{16}y - 3 = c + dy$$

1. Find μ_1, μ_2

$$\Rightarrow a = 3 = \mu_2 - \rho \frac{\sigma_2}{\sigma_1} \mu_1, \quad b = 4 = \rho \frac{\sigma_2}{\sigma_1}$$

$$\Rightarrow c = -3 = \mu_1 - \rho \frac{\sigma_1}{\sigma_2} \mu_2, \quad d = \frac{1}{16} = \rho \frac{\sigma_1}{\sigma_2}$$

$\Rightarrow$

→ substitute b on a $\Rightarrow u_2 - 4u_1 = 3$ --- ①

→ substitute d on c $\Rightarrow (u_1 - \frac{1}{16}u_2 = -3) \times 16$ --- ②

solve

$$u_2 - 4u_1 = 3$$

$$+ 16u_1 - u_2 = -48$$

$$12u_1 = -45 \Rightarrow$$

$$u_1 = -\frac{15}{4} = -3.75$$

$$\Rightarrow u_2 = 3 + 4u_1 = 3 + 4(-\frac{15}{4}) = -12$$

$$u_2 = -12$$

2. Find P :

$$\begin{pmatrix} P & z_1 \\ & z_2 \end{pmatrix} \begin{pmatrix} P & z_1 \\ & z_2 \end{pmatrix} = P^2$$

$$4 \times \frac{1}{16} = P^2$$

$$\frac{1}{4} = P^2 \Rightarrow P = \pm \frac{1}{2}$$

$P \neq -\frac{1}{2}$ why?

exp 3 : x

exp 4: let x, y have p.d.f

$$f(x, y) = \begin{cases} e^{-y} & , 0 < x < y < \infty \\ 0 & , \text{elsewhere} \end{cases}$$

calculate ρ

→ you can calculate ρ as before

$$\rho = \frac{E(xy) - E(x)E(y)}{\sqrt{E(x^2) - (E(x))^2} \sqrt{E(y^2) - (E(y))^2}}$$

$$\cdot E(x) = \int_0^{\infty} \int_0^y x f(x, y) dx dy = \int_0^{\infty} \int_0^y x e^{-y} dx dy$$

$$\cdot E(y) = \int_0^{\infty} \int_0^y y e^{-y} dx dy$$

$$\cdot E(x^2) = \int_0^{\infty} \int_0^y x^2 e^{-y} dx dy$$

$$\cdot E(y^2) = \int_0^{\infty} \int_0^y y^2 e^{-y} dx dy$$

$$\cdot E(xy) = \int_0^{\infty} \int_0^y xy e^{-y} dx dy$$

Def: let X_1, X_2 be Random variables with joint p.d.f $f(x_1, x_2)$ we defined the m.g.f of X_1, X_2 are follows:

$$M(t_1, t_2) = E(e^{t_1 X_1 + t_2 X_2}).$$

if expected value exists in a Region a Bound $(0,0)$.

Remark:

1. $M(0,0) = 1$ joint m.g.f X_1, X_2 at $(t_1, t_2) = (0,0)$.
2. $M(t_1, 0) = M_1(t_1)$ Marginal m.g.f of X_1
3. $M(0, t_2) = M_2(t_2)$ Marginal m.g.f of X_2 .
4. $\frac{\partial^{k+L}}{\partial t_1^k \partial t_2^L} M(t_1, t_2) \Big|_{(t_1, t_2) = (0,0)} = E(X_1^k, X_2^L) \quad k, L = 1, 2, \dots$

Back to exp 4:

lets find P using m.g.f

$$\begin{aligned} E(e^{t_1 X + t_2 Y}) &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{t_1 x + t_2 y} f(x, y) dx dy = \int_0^{\infty} \int_0^y e^{t_1 x + t_2 y} e^{-y} dx dy \\ &= \int_0^{\infty} \int_x^{\infty} e^{t_1 x + t_2 y} dy dx \end{aligned}$$

→

$$\int_0^{\infty} \int_0^y e^{t_1 x + t_2 y} e^{-y} dx dy = \int_0^{\infty} e^{t_2 y} e^{-y} \left(\int_0^y e^{t_1 x} dx \right) dy$$

$$= \int_0^{\infty} e^{(t_2-1)y} \left(\frac{e^{t_1 x}}{t_1} \Big|_0^y \right) dy$$

$$= \int_0^{\infty} e^{(t_2-1)y} \left(\frac{e^{t_1 y}}{t_1} - \frac{1}{t_1} \right) dy$$

$$= \int_0^{\infty} \frac{e^{(t_1+t_2-1)y}}{t_1} - \frac{e^{(t_2-1)y}}{t_1} dy$$

$$= \left(\frac{1}{t_1} \frac{e^{(t_1+t_2-1)y}}{t_1+t_2-1} - \frac{1}{t_1} \frac{e^{(t_2-1)y}}{t_2-1} \right) \Big|_0^{\infty}$$

$$= \left(\lim_{a \rightarrow \infty} \frac{1}{t_1} \frac{e^{(t_1+t_2-1)a}}{t_1+t_2-1} - \lim_{b \rightarrow \infty} \frac{1}{t_1} \frac{e^{(t_2-1)b}}{t_2-1} \right) - \left(\frac{1}{t_1(t_1+t_2-1)} - \frac{1}{t_1(t_2-1)} \right)$$

$$= - \frac{1}{t_1(t_1+t_2-1)} + \frac{1}{t_1(t_2-1)} \quad \text{assuming } t_1 \neq 0, t_1+t_2 < 1, t_2 < 1$$

$$= \frac{-(t_2-1) + (t_1+t_2-1)}{t_1(t_2-1)(t_1+t_2-1)}$$

$$= \frac{-t_2 + 1 + t_1 + t_2 - 1}{t_1(t_2-1)(t_1+t_2-1)}$$

$$= \frac{1}{(1-t_1-t_2)(1-t_2)}, \quad t_2 < 1, t_1+t_2 < 1$$

is the double
dy dx

$$E(x) = ?$$

$$\cdot M_1(t_1) = \frac{1}{1-t_1} = (1-t_1)^{-1}, \quad t_1 < 1$$

$t=0$

$$\cdot \bar{M}_1(t_1) = (1-t_1)^{-2} \Rightarrow E(x) = 1 = \mu_1 = \mu_x$$

$$\cdot \bar{M}_2(t_1) = 2(1-t_1)^{-3} \Rightarrow E(x^2) = 2 = \sigma_x^2 = \sigma_1^2$$

$$\cdot M_2(t_2) = \frac{1}{(1-t_2)^2} = (1-t_2)^{-2}, \quad t_2 < 1$$

$$\cdot \bar{M}_2(t_2) = 2(1-t_2)^{-3} \Rightarrow E(y) = 2 = \mu_2 = \mu_y$$

$$\cdot \bar{M}_2(t_2) = 6(1-t_2)^{-4} \Rightarrow E(y^2) = 6 = \sigma_y^2 = \sigma_2^2$$

$$\cdot \frac{\partial^2 M(t_1, t_2)}{\partial t_1 \partial t_2} \bigg|_{(t_1, t_2) = (0,0)} = 3 \Rightarrow E(x, y) = 3$$

$$\Rightarrow \rho = \frac{E(xy) - E(x)E(y)}{\sigma_x \sigma_y} = \frac{3 - (1)(2)}{\sqrt{1} \sqrt{2}} = \frac{1}{\sqrt{2}}$$

□